

# Dynamic Factor Models in Real Time: A New Approach and an Implementation in Matlab

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Moderní nástroje pro finanční analýzu a modelování  
Praha, 5 June 2014

# Outline of the presentation

1. Motivation
2. My contribution
3. Implementation in Matlab
  - ▶ Simulation Study

# Motivation

## Stock and Watson (2011)

The premise of dynamic factor models (DFM) is that a few latent dynamic factors  $f_t$  drive the comovements of high-dimensional vector of time-series variables  $X_t$ , which is also affected by a vector of disturbances  $\varepsilon_t$ :

$$X_t = \underbrace{\sum_i \Lambda_i f_{t-i}}_{\equiv \chi_t = \text{common component}} + \underbrace{\varepsilon_t}_{\text{idiosyncratic part}} .$$

# Applications

Dynamic factor models have a wide range of applications:

- ▶ Near-term forecasting
- ▶ Analysis of macroeconomic comovements
- ▶ Analysis of commodity markets
- ▶ Finance: yield curve
- ▶ and more ....

Stock and Watson (2011) distinguish several generations of DFM

## First generation

First generation DFM have been estimated in time-domain using MLE

$$X_t = \Lambda_0 f_t + e_t$$
$$f_t = \Psi_1 f_{t-1} + \dots + \Psi_j f_{t-j} + G\eta_t,$$

By the **parametric** specification of distribution for  $\eta_t$  (Gaussian iid errors) and  $e_t$  (usually Gaussian stationary process), the model can be converted to a state space form and the likelihood function evaluated using the Kalman filter:

- ▶ static nature:
  - ▶ no lead-lag relation among  $X_t$
  - ▶ not always plausible, e.g. the unemployment cycles lags the output cycles in most advanced countries (Brůha & Polanský, 2014; Andrlé, Brůha, Solmaz 2013, 2014).

## Second generations /a

### Non-parametric 'averaging' errors:

- ▶ (Generalized) principal component estimation
- ▶  $\min_{f_1, \dots, f_T, \Lambda} T^{-1} \sum_{t=1}^T (X_t - \Lambda f_t) \Sigma_e^{-1} (X_t - \Lambda f_t)'$ ,
- ▶ Leads to the eigenvalue decomposition of  $\Sigma_e^{-1/2} \widehat{\Sigma}_X \Sigma_e^{-1/2}$ ,
- ▶ Weak assumptions on dgp for  $f_t$
- ▶ Standard PCA if  $\Sigma_e$  is replaced by the identity matrix (Bai & Ng, 2002, 2003)
- ▶ Hard to estimate precisely  $\Sigma_e$ 
  - ▶ Forni, Halli, Lippi and Reichlin (2005) suggest estimating it by non-parametric spectral density (a dynamic model)
- ▶ Again **static** model, even if  $\Sigma_e$  estimated from a dynamic model

## Second generations /b

### **Dynamic** principal component analysis:

- ▶ Based on the frequency-domain PCA by Brillinger (1964)
- ▶ Introduced to empirical economics by Forni, Halli, Lippi and Reichlin (2000)
- ▶ A non-parametric method:
  - ▶ weak assumption on data generating process (rational spectral density),
  - ▶ the spectral density can be estimated by a non-parametric (Bartlett) approach.
- ▶ The spectral density is subjected to PCA and then transformed back to time-domain filter using a two-sided linear filter

## Second generations /c

The transformation to time domain leads to a two-sided filter:

$$\chi_t = \sum_{i=-K}^K \omega_k X_{t-k},$$

where  $\chi_t$  is the common component (driven by the common factors), and  $\omega_k$  are filtered weights.

Implications:

- ▶ Forni, Halli, Lippi and Reichlin (2000) call it 'two-sided' DFM
  - ▶ versus the 'static' one-sided DFM based on generalized PCA by Forni, Halli, Lippi and Reichlin (2005) outlined above
- ▶ The common component  $\chi_t$  cannot be estimated at the beginning and at the end of the sample
- ▶ Problems for many applications (real-time data)
- ▶ **My contribution: how to estimate the common component on the whole sample.**

## Third generation – state space models

They can be applied to a variety of **parametric** models using sophisticated tools of modern econometrics, such as

- ▶ the EM algorithm (Bańbura & Modugno, 2010),
- ▶ the Gibbs sampler (Bai & Wang, 2012),
- ▶ the two-step estimation (Doz, Giannone and Reichlin, 2006).

These models:

- ▶ are truly dynamic (unlike the first generations and one-sided PCA),
- ▶ can accommodate various issues (missing data, asynchronous data release, ...),
- ▶ the common component estimated on the *whole* sample,
- ▶ but **are parametric**.

# Quest

## Quest

- ▶ Is it possible to propose:
  - ▶ a *genuinely* dynamic model,
  - ▶ based on non-parametric (*frequency-domain*) approach,
  - ▶ with the common component estimable on *the whole sample*,
  - ▶ and possibly accommodating *other issues*, such as asynchronous data releases?

Yes

... and this is my contribution.

## How is it done?

Start with the common component representation:

$$X_t = \underbrace{\chi_t}_{\equiv \sum_i \Lambda_i f_{t-i}} + \varepsilon_t,$$

Now consider the linear projection of  $\chi_t$  on  $X_S \equiv \{X_s\}_{s \in S}$ , where  $S$  is an *arbitrary* index set. The linear projection is given as:

$$\mathbb{E}^* [\chi_t | X_S] = \mathbb{E} [\chi_t X_S^T] \mathbb{E} [X_S X_S^T]^{-1} X_S,$$

Matrices such as  $\mathbb{E} [\chi_t X_S^T]$  can be derived from the spectral density of the data:

- ▶ as in 'second-generation' DFM (DPCA)
- ▶ hence, frequency-domain DPCA meets the linear regression.

# How to get the common component on the whole sample?

The adaptation of the index set  $S$ :

- ▶ The index set can be adapted to the beginning and end of the sample (hence, real time data).
  - ▶ assume you wish to get  $\mathbb{E}^* [\chi_t | \{X_{t-1}, X_t, X_{t+1}\}]$
  - ▶ trivial in the middle of the sample,
  - ▶ replace it by  $\mathbb{E}^* [\chi_T | \{X_{T-1}, X_T\}]$  at the end of the sample!
- ▶ The index set can be adapted for missing data (e.g. due to asynchronous data release).

# Finite-sample properties

The estimation of covariance matrices  $\mathbb{E} [\chi_t \chi_S^T]$  and  $\mathbb{E} [X_S X_S^T]$  can be imprecise in finite samples.

How to overcome this complication:

- ▶ Use robust regression in computing the projection  $\mathbb{E}^* [\chi_t | X_S]$ !
- ▶ I use a ridge regression with good properties.

## Matlab codes

I build a set of Matlab routines that can be used to run a set of non-parametric DFM

`sfactor.m` computes the static PCA

`dfactor.m` computes the 'two-sided' version of the DPCA

- ▶ for particular choices, equivalent to the 'one-sided' DFM representation,

`tfactor.m` implements the approach outlined in this presentation.

I plan to put these routines on Matlab Exchange File soon

## A simulation study /1

Forni, Halli, Lippi and Reichlin (2005) proposed four computational experiments with DFM.

**M1** read as

$$x_{it} = \lambda_i f_t + \alpha c_i \epsilon_{it} \text{ with } (1 - .5L)f_t = u_t,$$

and the shocks  $\epsilon_{it}$ ,  $u_t$ ,  $\lambda_i$  are independent standard normal variables,  $c_i \sim \mathcal{U}_{[0.1 \ 1.1]}$  and  $\alpha$  is calibrated so that the average idiosyncratic-common variance ratio is 1.

**M2** reads as:

$$x_{it} = \sum_{k=0}^3 a_{ik} u_{1,t-k} + \sum_{k=0}^3 b_{ik} u_{2,t-k} + \alpha c_i \epsilon_{it},$$

where  $a_{ik}$ ,  $b_{ik}$  and the shocks are standard normal variables and  $c_i$  and  $\alpha$  are as above.

## A simulation study /2

**M3** reads as:

$$x_{it} = \sum_{k=l_i}^{l_i+2} \lambda_{k-l_i, i} f_{t-k} + \xi_{it},$$

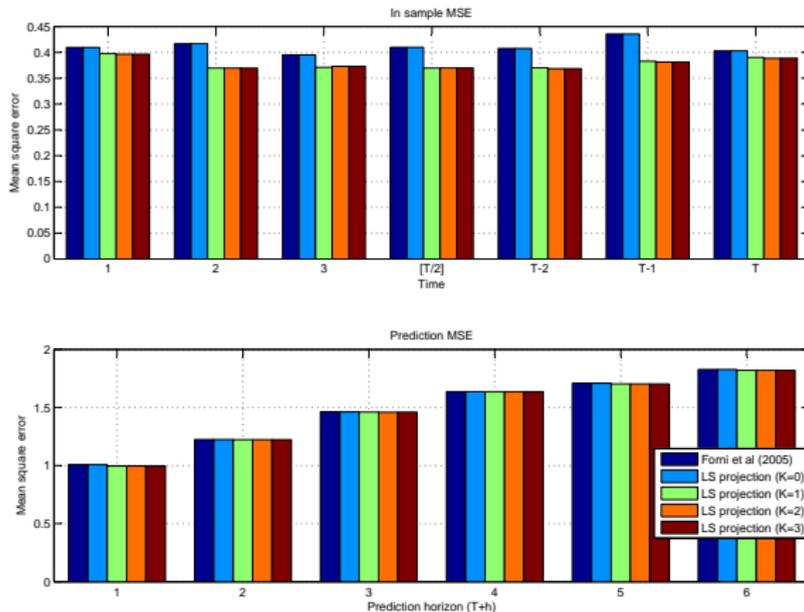
where  $(1 - .5L)f_t = u_t$ ,  $\xi_{it} = \alpha c_i(\varepsilon_{it} + \varepsilon_{i+1t})$ ,  $u_t$  and  $\varepsilon_{it}$  are independent standardized normal random variables,

$$l_i = \begin{cases} 0 & \text{for } i \leq m \\ 1 & \text{for } i \in \{m+1, \dots, 2m\} \\ 2 & \text{for } i \geq 2m+1 \end{cases}$$

**M4** is the same as **M3**, but without cross-sectional dependence ( $\xi_{it} = \alpha c_i \varepsilon_{it}$ ).

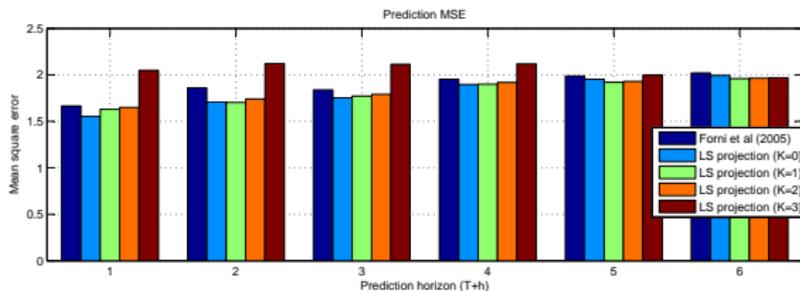
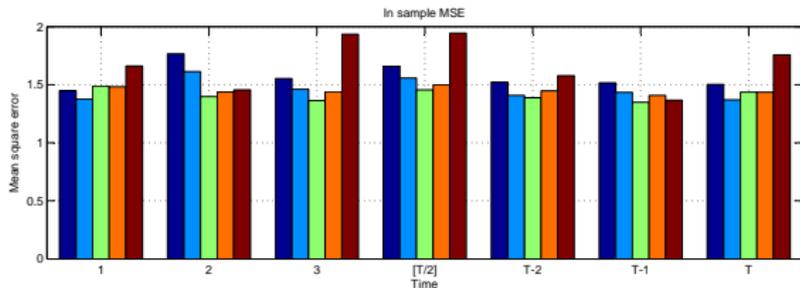
# Model M1 – known covariance matrices

Marginal efficiency gains for this model



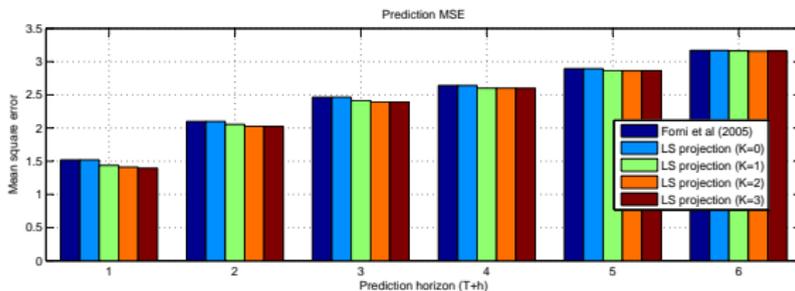
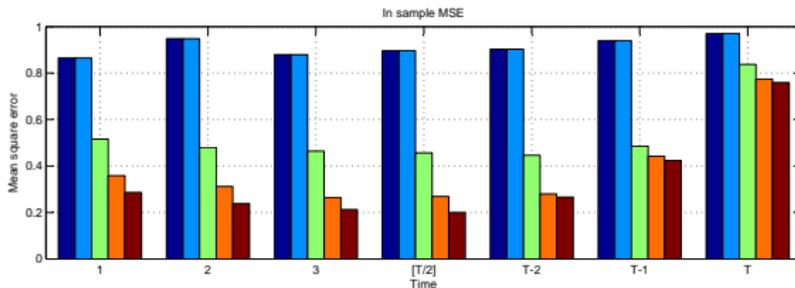
# Model M1 – unknown covariance matrices

Gains disappear if covariance matrices are to be estimated



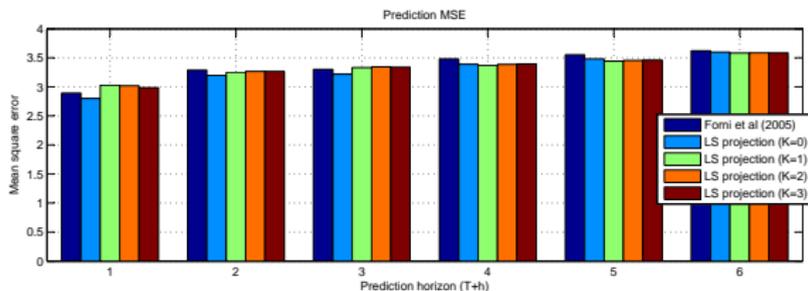
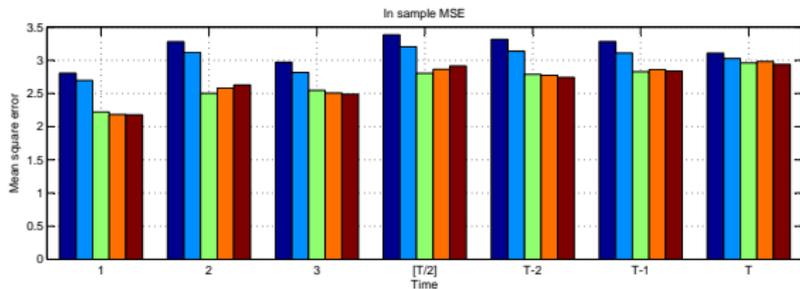
# Model M3 – known covariance matrices

Large efficiency gains at the beginning of the sample



# Model M3 – unknown covariance matrices

The efficiency gains survive even in small samples ( $T=50$ ,  $N=20$ )



# Disclaimer

The tool presented here was partly developed under the CNB research project B1/10. Nevertheless, the views expressed here are mine and do not necessarily reflect the position of the Czech National Bank.

I thank to Michal Andrlé for comments and encouragement. However, any errors are solely my own responsibility.

## Concluding slide

Thank you for your attention.

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